

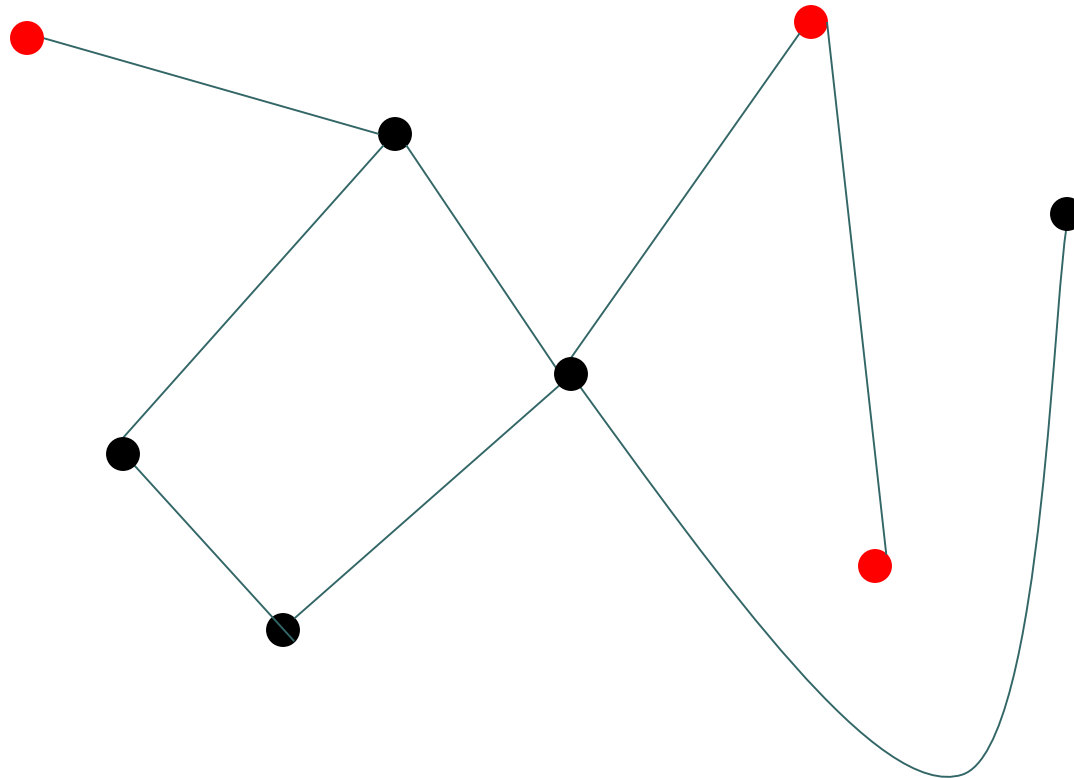
MEAN FIELD CONDITIONS FOR COALESCING RANDOM WALKS

Roberto Imbuzeiro Oliveira
IMPA, Rio de Janeiro

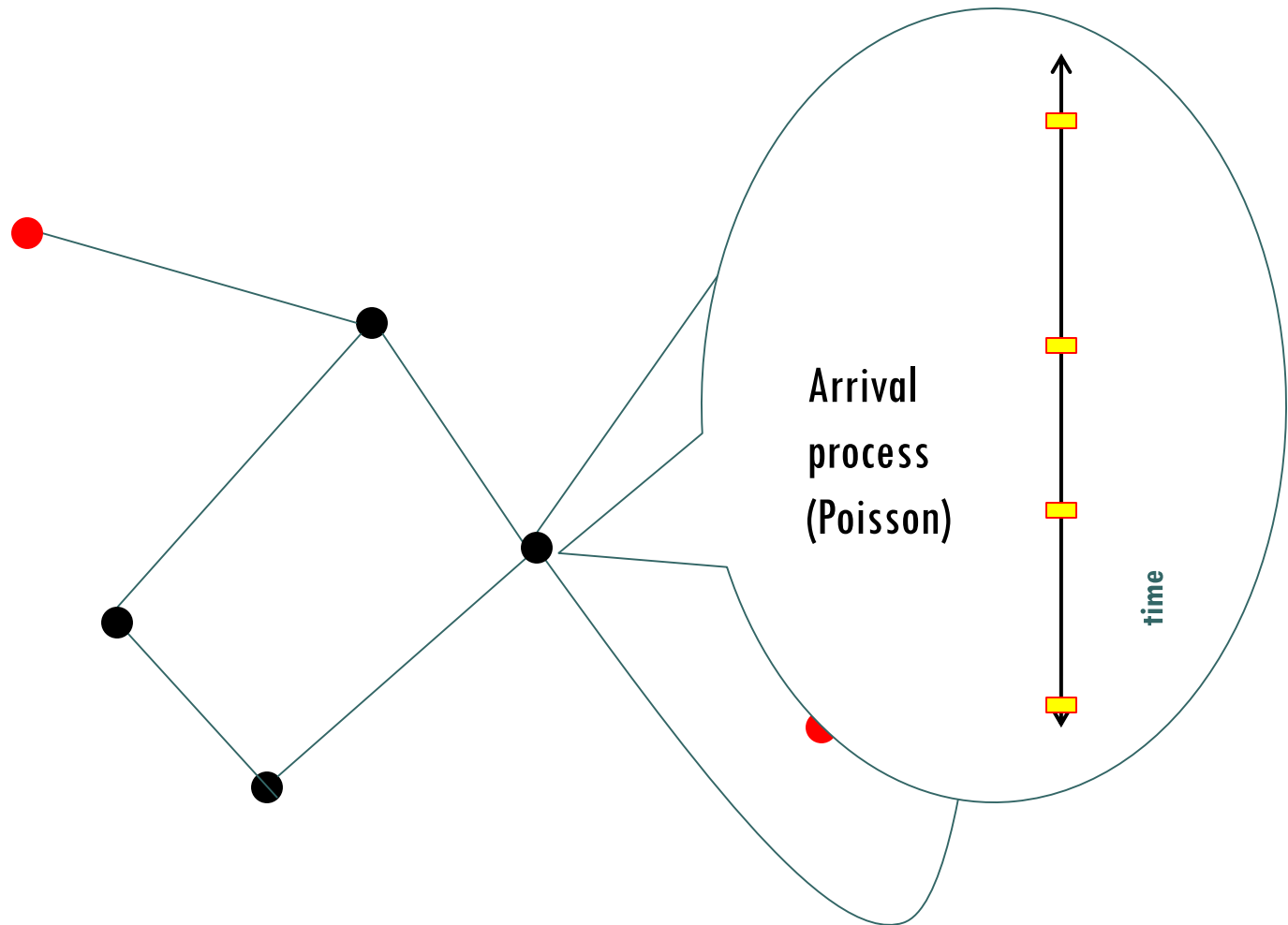
SPA 2014 (Buenos Aires)

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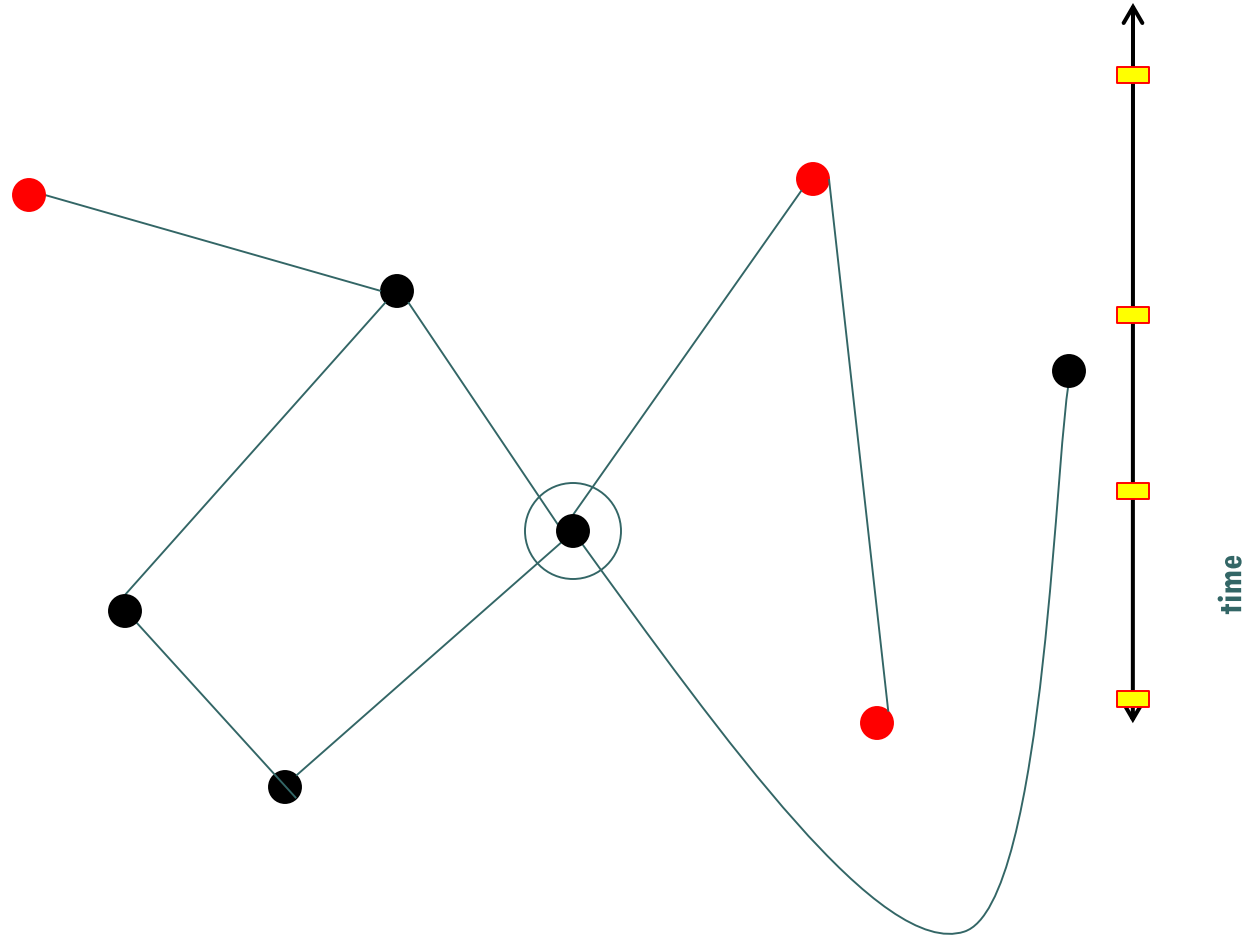
WHAT IS THE VOTER MODEL?



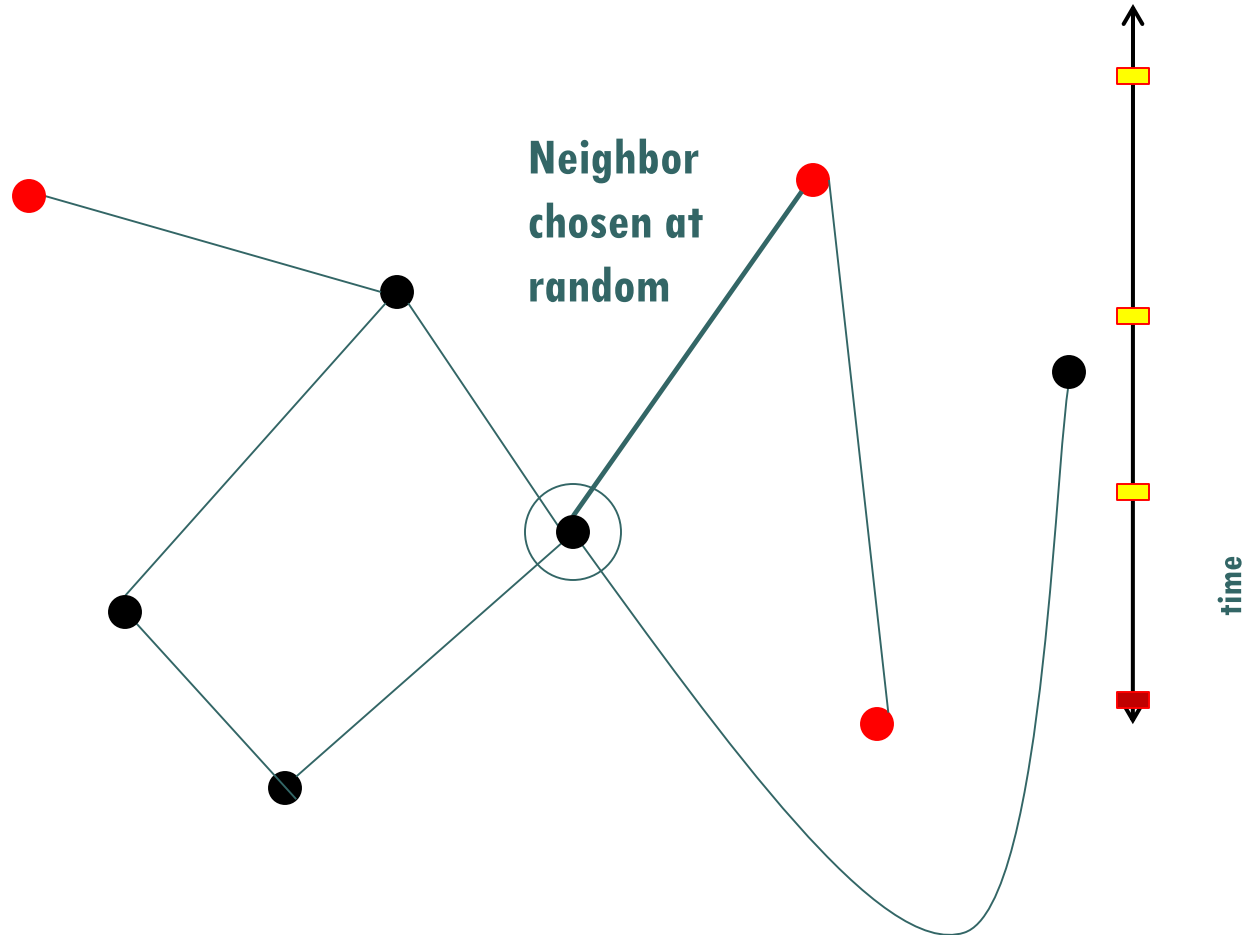
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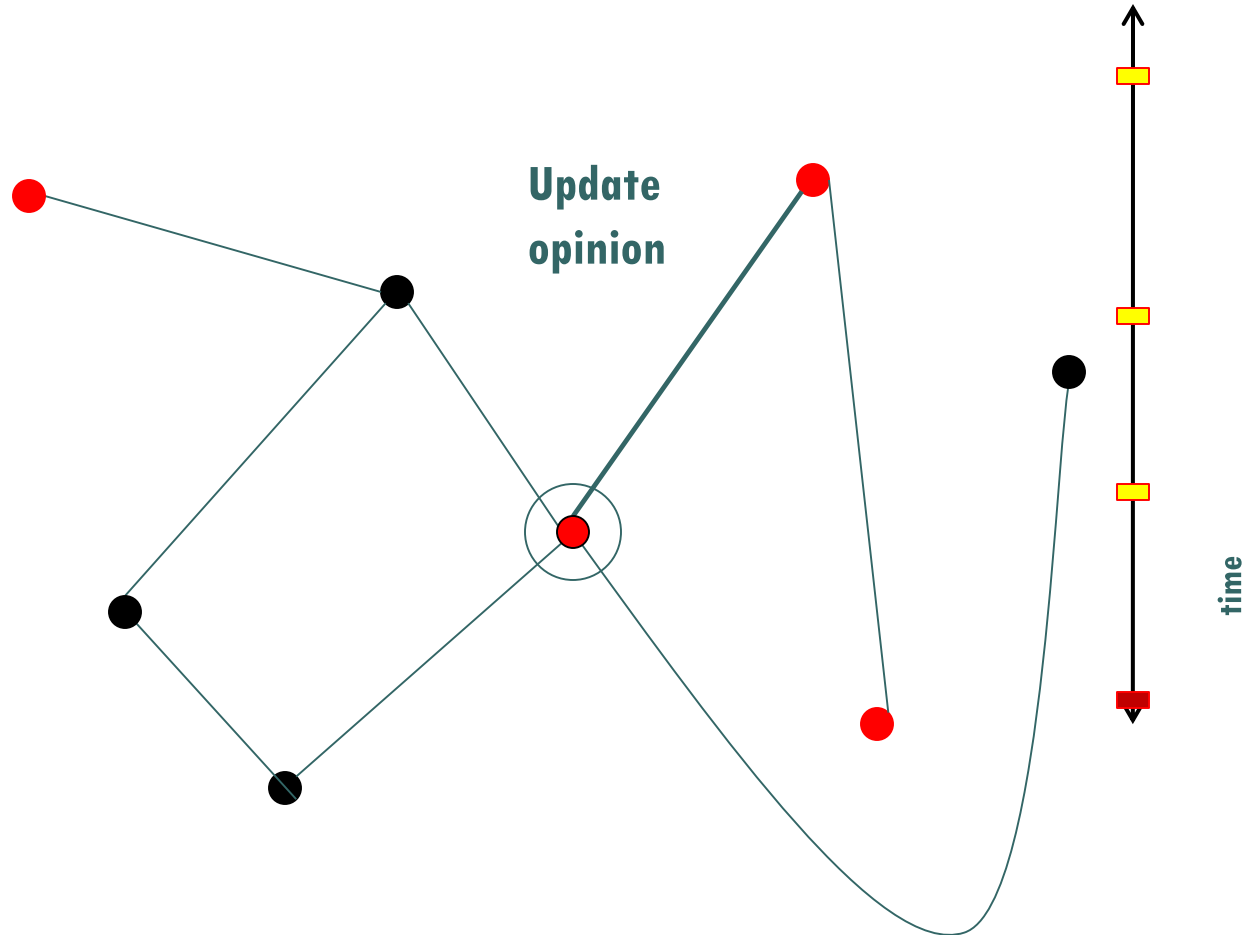
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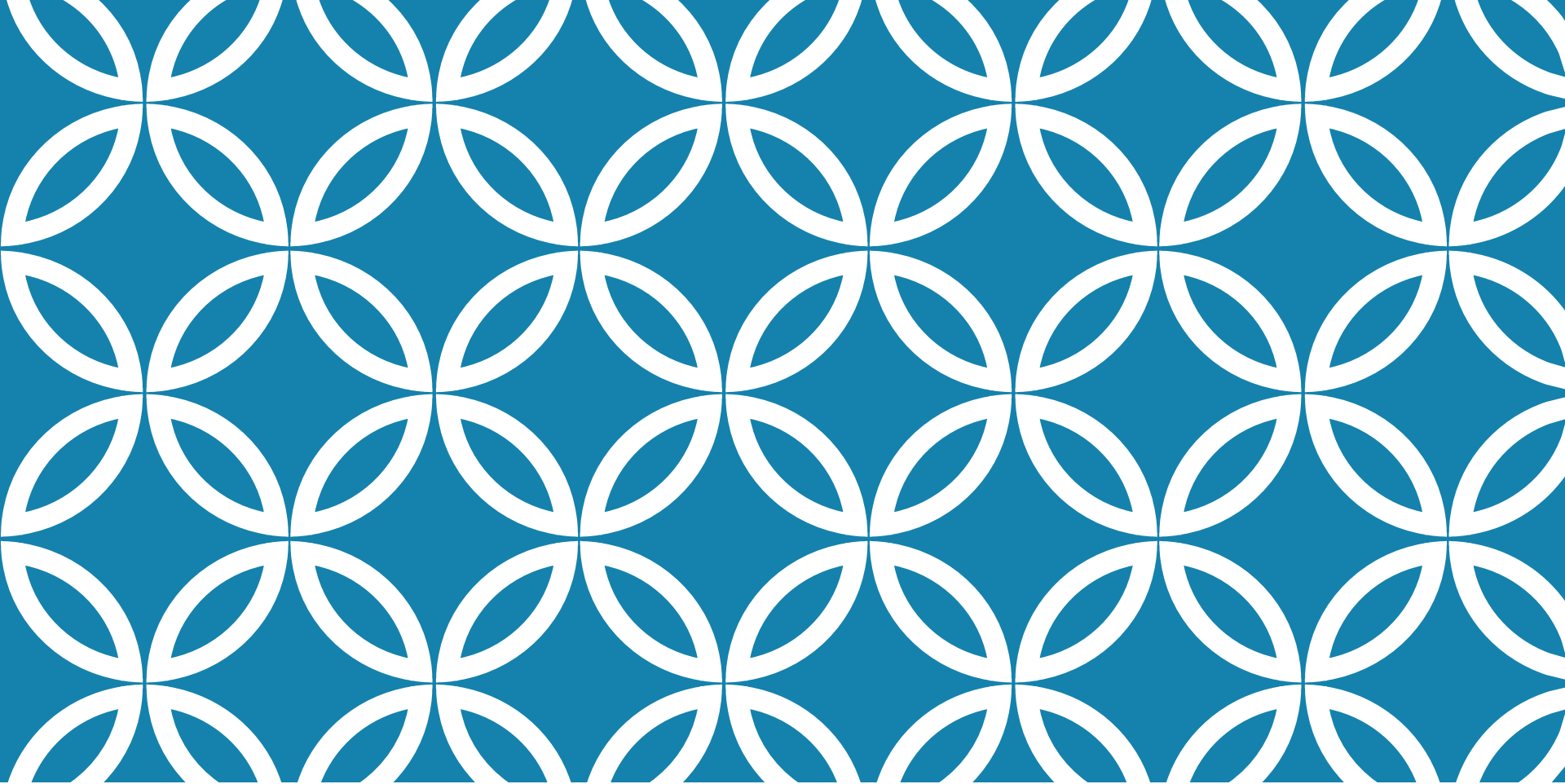


WHAT IS THE VOTER MODEL?



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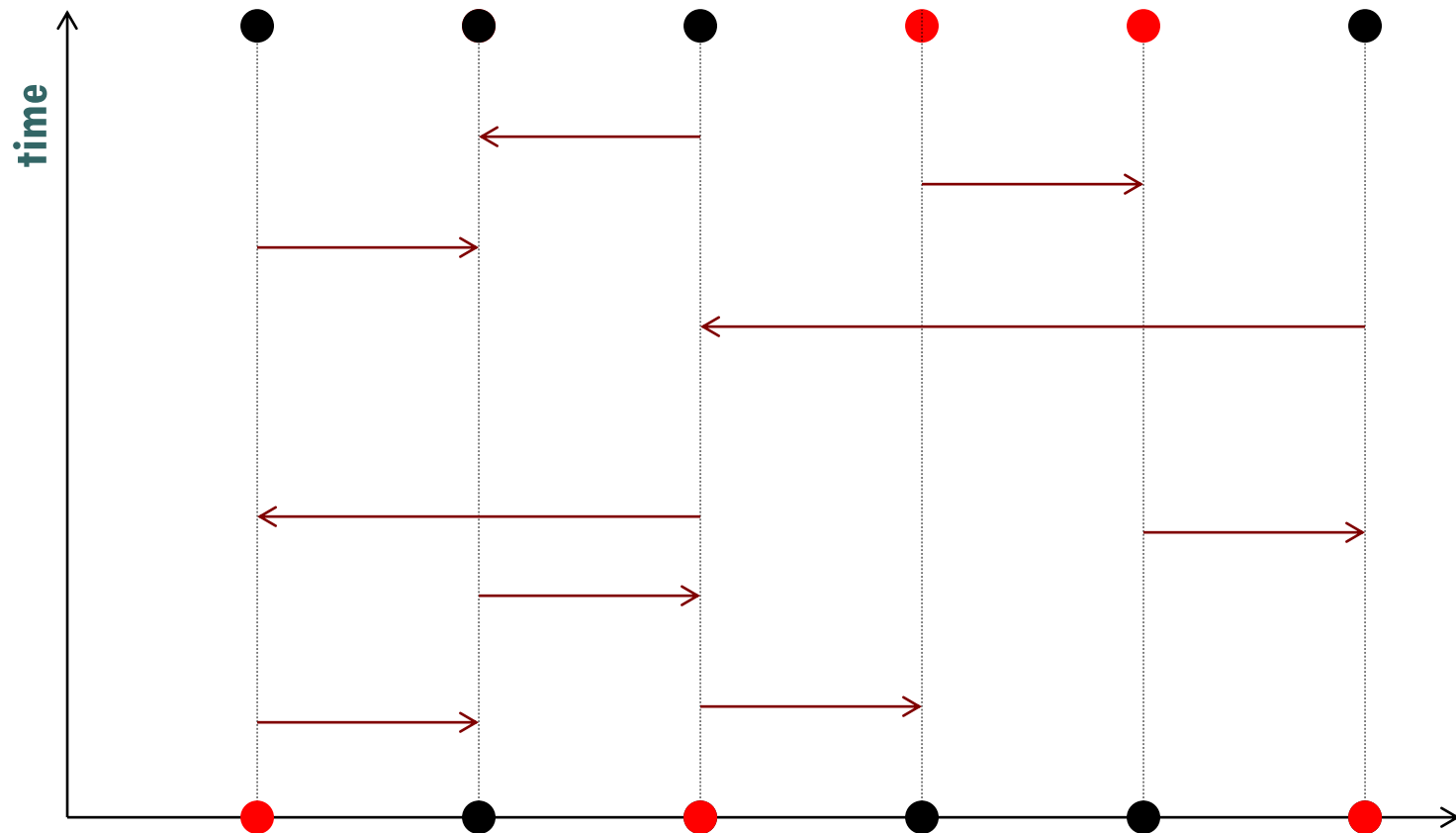




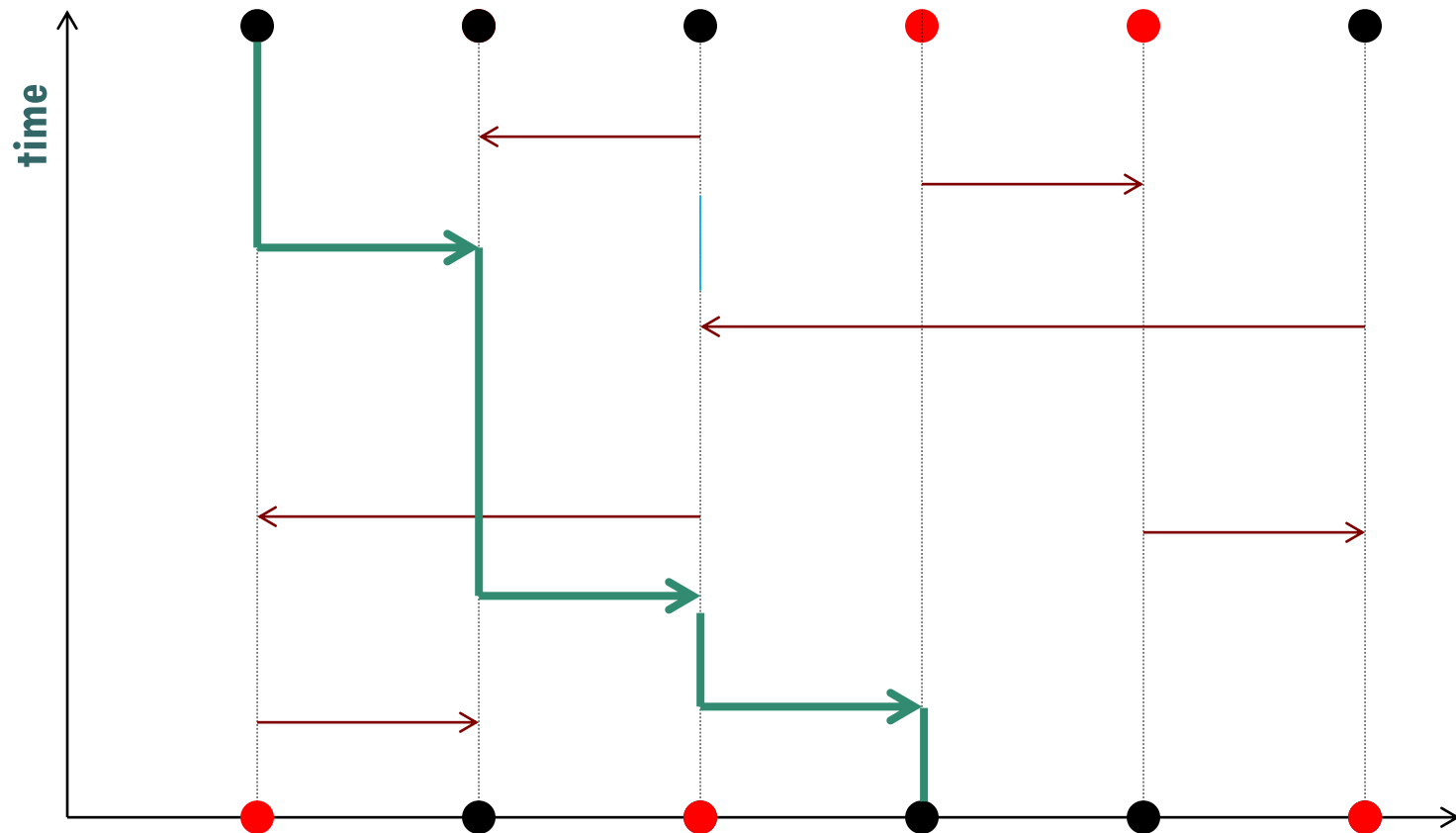
VOTERS, RANDOM WALKERS, AND DUALITY

The original motivation for studying coalescing random walks is the voter model.

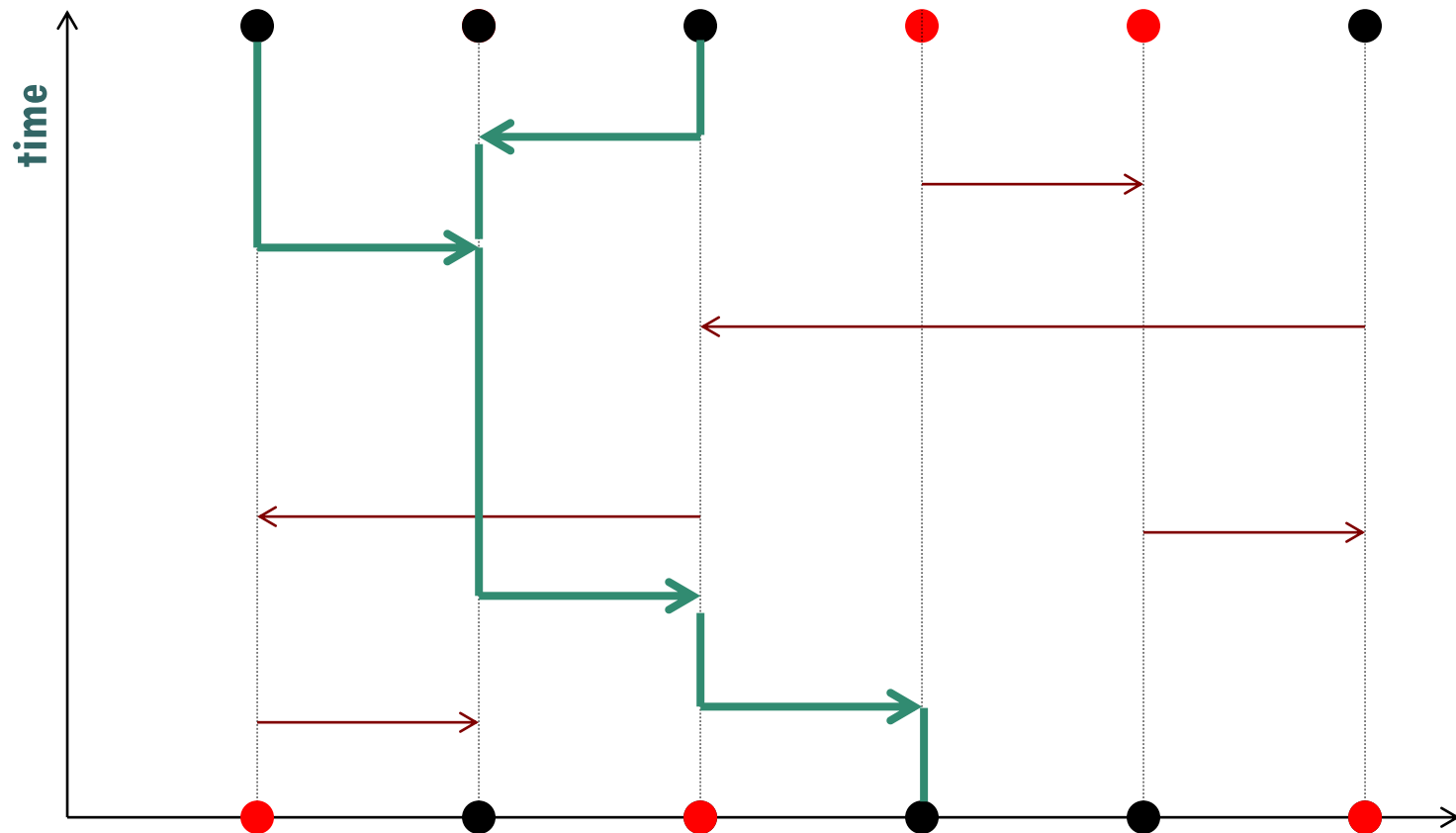
DUALITY



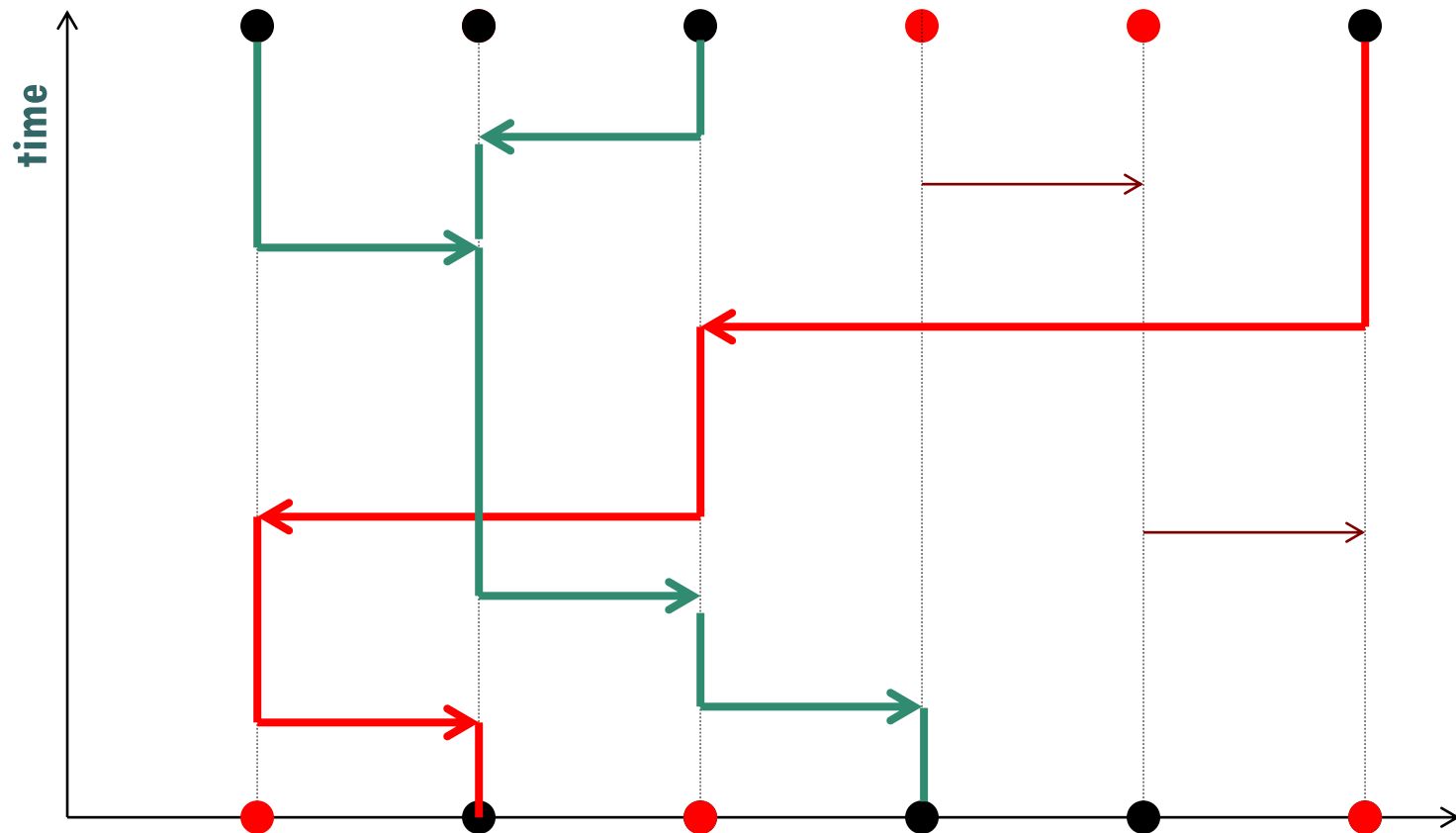
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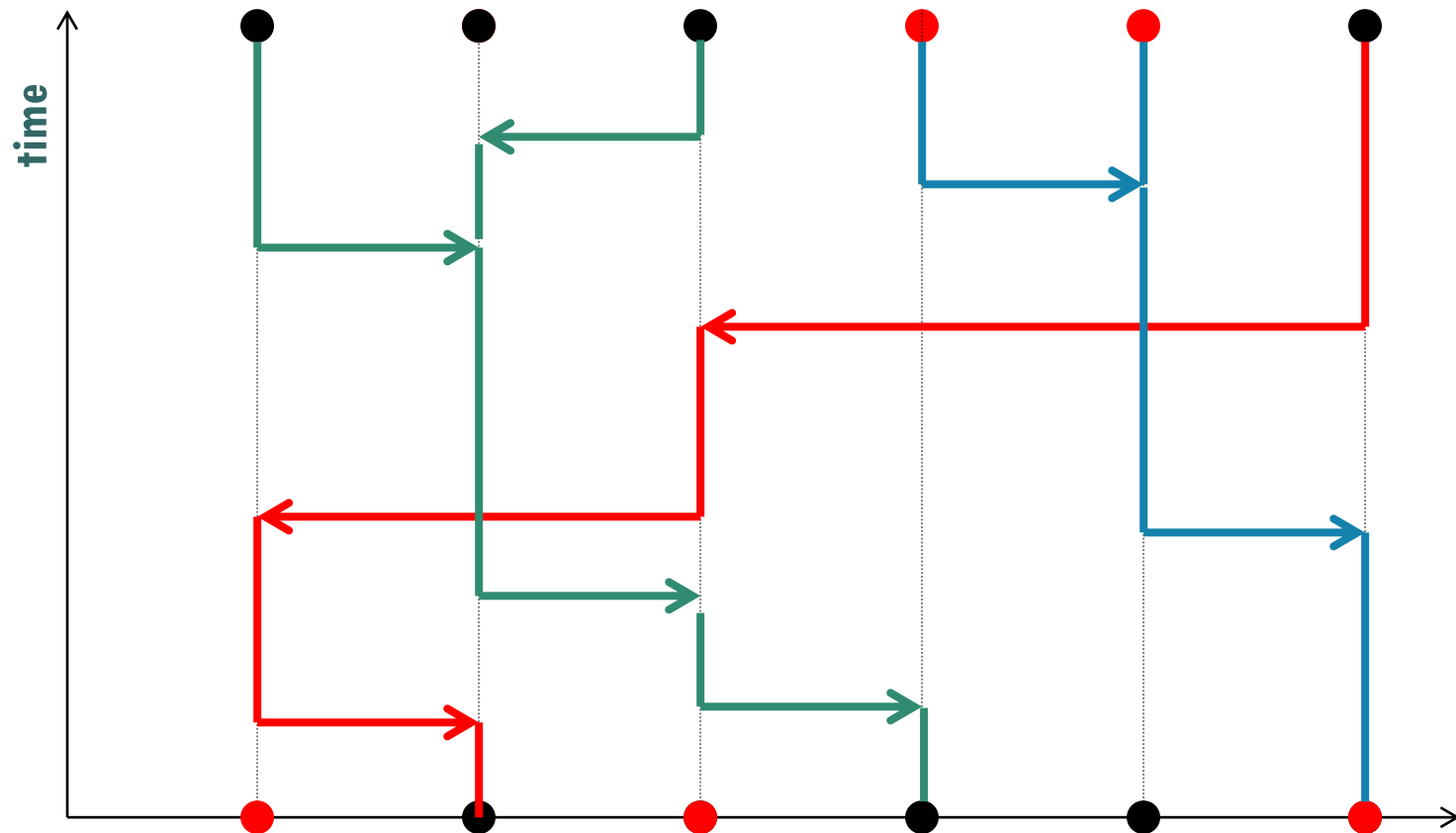
DUALITY



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DUALITY

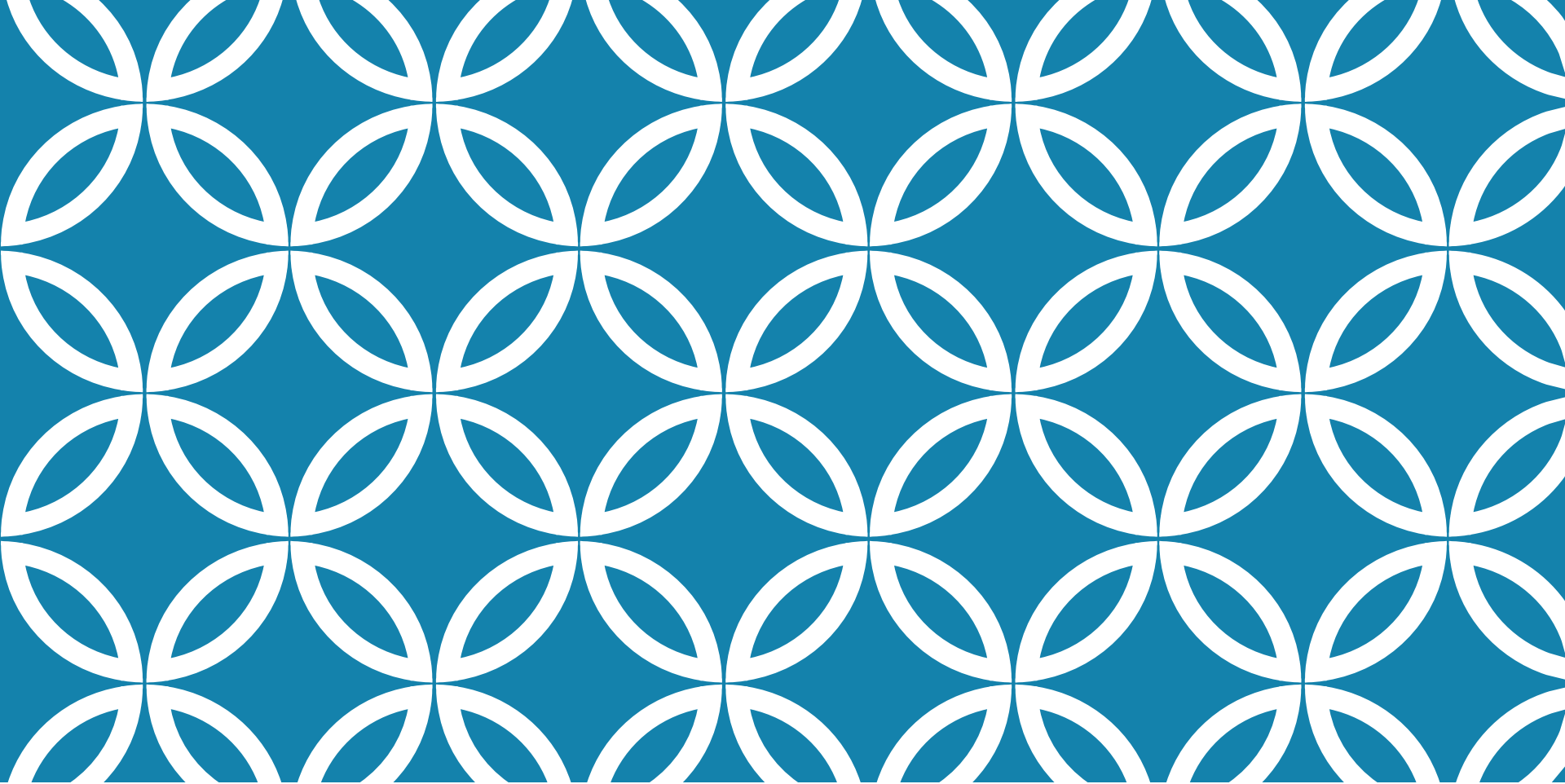
The upshot is that the voter model is dual to

Coalescing random walks,

the main subject of this talk. Will discuss:

C := full coalescence time.

Results extend to voters with i.i.d. initial opinions.



MEAN FIELD BEHAVIOR FOR FULL COALESCENCE

The case of the complete graph is easy. Other cases turn out to be similar.

THE COMPLETE GRAPH

System of n coalescing random walks on a complete graph K_n .
(ie. all transition rates equal to 1).

THE COMPLETE GRAPH

System of n coalescing random walks on a complete graph K_n .

(ie. all transition rates equal to 2).

Time to move from k to $k-1$ particles:

$$\mathbb{P}(Z_k \geq t) = e^{-\binom{k}{2} t}$$

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$$\mathbb{P}(Z_k \geq t) = e^{-\binom{k}{2} t}$$

ie. exponential with mean $\binom{k}{2}^{-1}$.

THE COMPLETE GRAPH

System of n coalescing random walks on a complete graph K_n .

(ie. all transition rates equal to 2).

$$C = \sum_{k=2}^n Z_k$$

$\{Z_k\}_n$ independent, $Z_n = d \exp\left(\frac{1}{\binom{n}{2}}\right)$.

THE COMPLETE GRAPH

System of n coalescing random walks on a complete graph K_n .

(ie. all transition rates equal to 2).

$$C \stackrel{\sim}{=} \sum_{k=2}^{\infty} z_k$$

$$\mathbb{E}(C) \sim 2 = 2 \mathbb{E}(\text{Meet of 2})$$

A RESULT BY COX

Cox'91: Consider CRW based on simple random walk in

$$(\mathbf{Z}_n)^d$$

where **d is at least 2**. Then:

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A RESULT BY COX

Cox'91: Consider CRW based on simple random walk in

$$(\mathbb{Z}_n)^d$$

where d is at least 2. Then:

$$\subseteq \mathcal{M}_{\mathbb{Z}_n^d}$$

$$n \rightarrow \infty \Rightarrow$$

$$\sum_{k \geq 2}$$

$$Z_k$$

indep.
exponentials,
means
 $= \left(\frac{1}{2}\right)^{-1}$

$\mathbb{E}(\text{Meeting time of } Z)$

A PROBLEM FROM ALDOUS AND FILL

Prove that this is universal over large transitive graphs
with **relaxation time** small.

A PROBLEM FROM ALDOUS AND FILL

PROBLEM:

$$\frac{C}{M_{G_n}}$$

$n \rightarrow \infty \Rightarrow$

$$\sum_{k \geq 2}$$

$$Z_k$$

indep.
exponentials
mean
 $\left(\frac{k}{2}\right)^{-1}$

$$E(\text{Meeting time of } Z) \gg t_{rel}^{G_n}$$

ASSUMPTION

A PROBLEM FROM ALDOUS AND FILL

Prove that this is universal over large transitive graphs
with **relaxation time** smaller than expected meeting time.

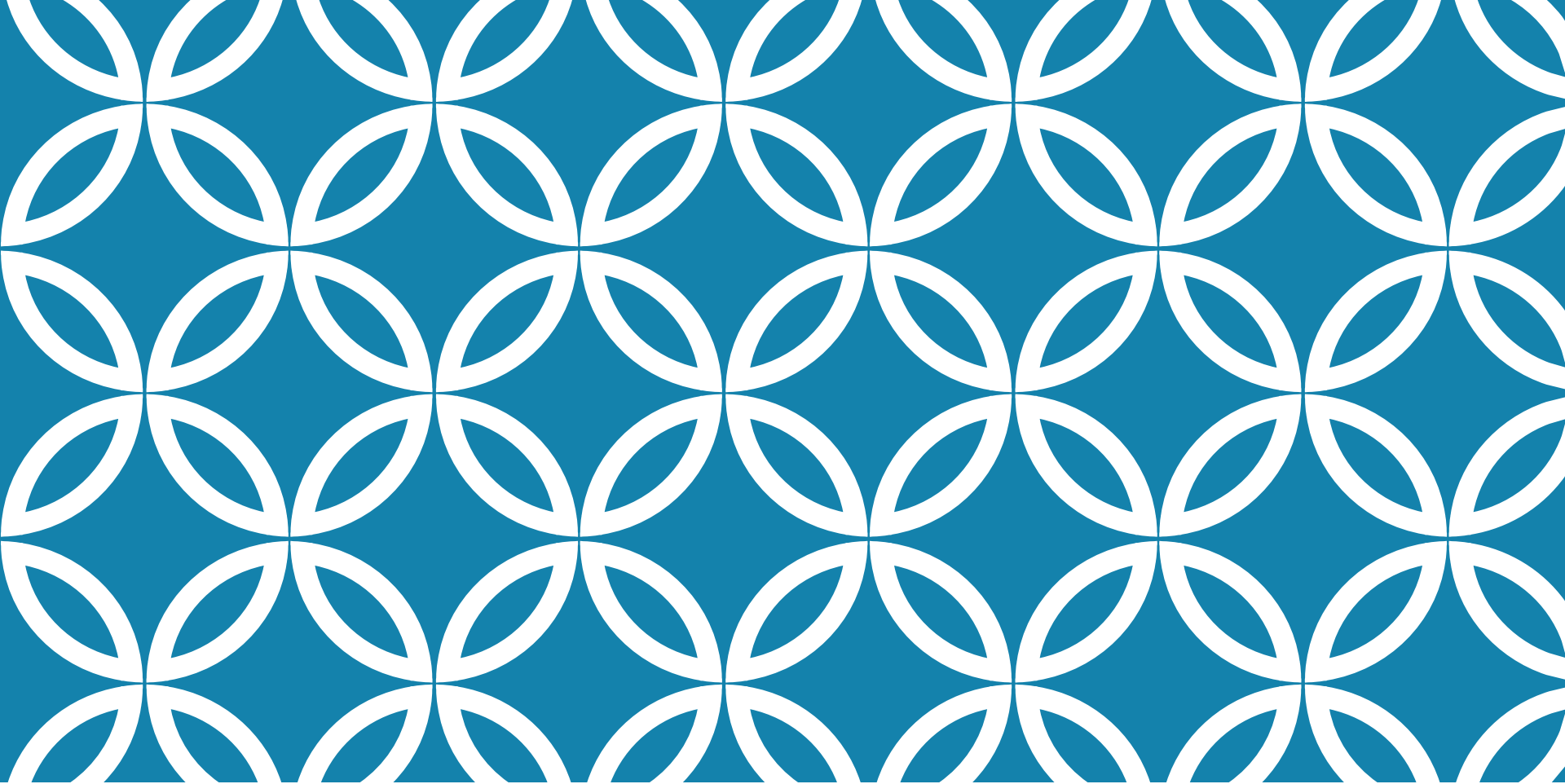
Some assumption is needed: no mean field behavior for star graphs
or one-dimensional cycles.

A MORE GENERAL PROBLEM BY DURRETT

In *Random Graph Dynamics* Durrett studies the same kind of problem over certain **random graphs**.

Those have **power law degrees** and are “very non transitive” in many ways.

Nevertheless, D. obtains some partial results in the direction of **universality of mean field behavior**.



MAIN RESULTS

Mean field behavior is indeed very general. We give two results.

A THEOREM FOR TRANSITIVE GRAPHS

$\{Q_n\}_n$: sequence of reversible,
transitive chains on
finite spaces.

$\{\mu_n\}_n$: expected meeting time
of 2 indep. Q_n -walks

A THEOREM FOR TRANSITIVE GRAPHS

$\{Q_n\}_n$

: sequence of reversible,
transitive chains on
finite spaces.

$t_{\text{mix},n}$
= mixing time

μ_{min}

expected meeting time
of 2 indep. Q_n -walks

A THEOREM FOR TRANSITIVE CHAINS

$$\frac{C_n}{M_n} \xRightarrow{n \rightarrow +\infty} \sum_{k \geq 2} z_k$$

whenever $\frac{t_{\text{mix}, n}}{M_n} \xrightarrow{n \rightarrow +\infty} 0$

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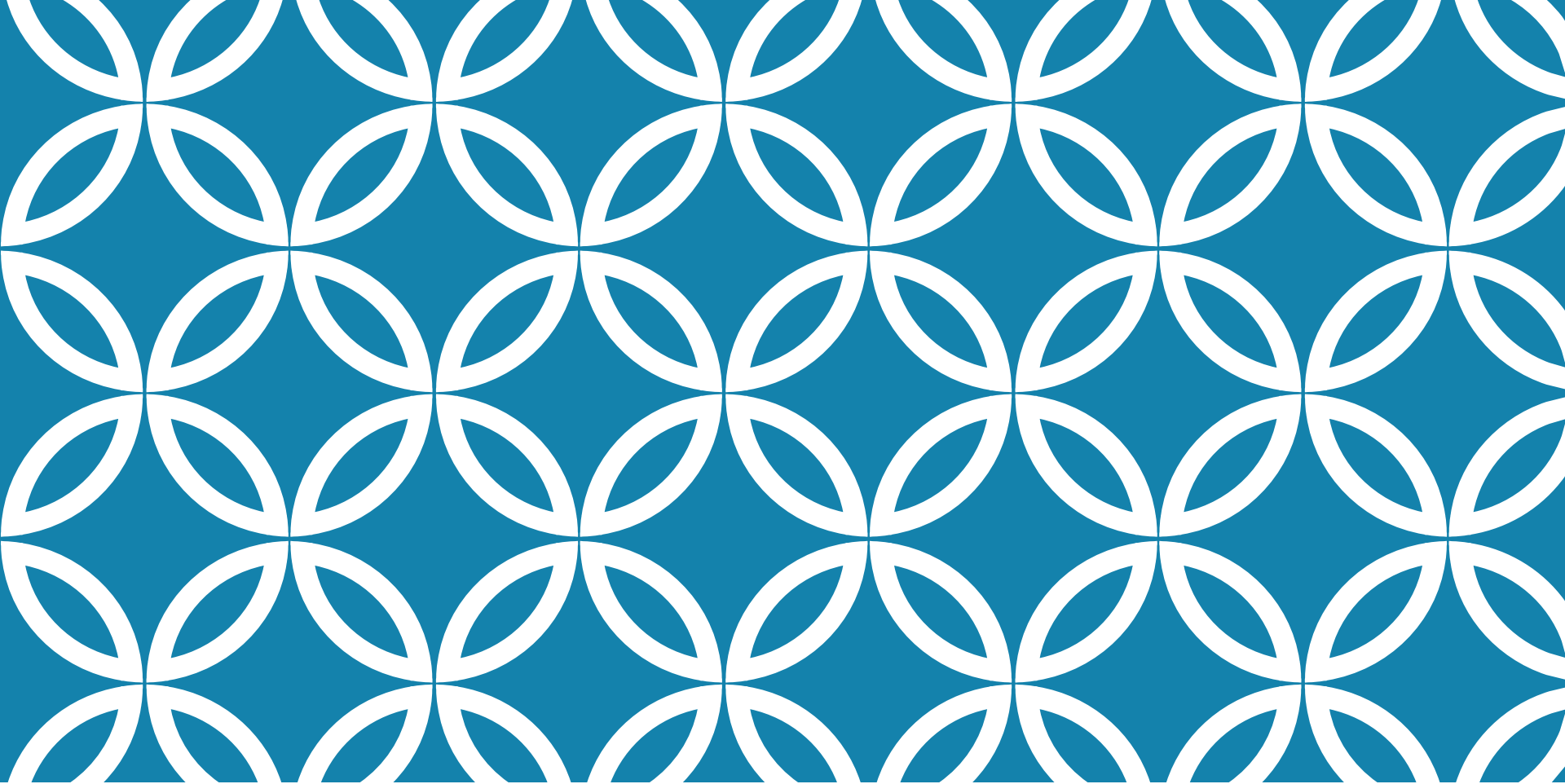
[Aldous/Fill: $t_{\text{rel},n}/M_n \rightarrow 0$.]

A THEOREM FOR GENERAL CHAINS

We also have a theorem not requiring transitivity or reversibility, with messier assumptions.

It covers the random graphs of Durrett + many other examples (eg. supercritical percolation in 3 or more dimensions).

There certainly is room for improvement here.



MAIN PROOF IDEAS

Exponential hitting times,
with good error bounds +
quantiles + control of big
bang phase.

THE THEOREM FOR TRANSITIVE CHAINS

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whenever $\frac{t_{\text{mix}, n}}{M_n} \xrightarrow{n \rightarrow +\infty} 0$

COMPARE WITH COMPLETE GRAPH

The random variables Z_k have a clearly defined meaning in the complete graph case.

Time to move from k to $k-1$ particles:

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BASIC IDEA: prove similar result for general chains.

CONNECTION WITH HITTING TIMES

k random walks
on V

\Rightarrow

one random
walk on
 V^k (product)

CONNECTION WITH HITTING TIMES

k random walks
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\Rightarrow

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 V^k (product)

$T_k = 1^{\text{st}}$ coalescence
among k

\Rightarrow

Hitting time of

$$\Delta_k = \{ (x_i)_{i=1}^k : \exists 1 \leq i < j \leq k \text{ s.t. } x_i = x_j \}$$

EXPONENTIAL HITTING TIMES

Well-known “metatheorem” (Aldous, Aldous/Brown,...).

$P = \text{chain on } \Omega$, $A \subset \Omega$ with
 π stationary, $\mathbb{E}_{\pi}(\tau_A) \gg t_{\text{mix}}^P$.

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$$\mathbb{P}_\pi(\tau_A / \mathbb{E}_\pi(\tau_A) > t) \approx e^{-t}$$

EXPONENTIAL HITTING TIMES

T_k = 1st coalescence among k \Rightarrow Hitting time of $\Delta_k = \{(x_i)_{i=1}^k : \exists 1 \leq i < j \leq k, x_i = x_j\}$

$$(k=2) \Rightarrow \mathbb{E}_\pi(T_2) = M_n \gg t_{\text{mix}}$$

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EXPONENTIAL HITTING TIMES

T_k = 1st coalescence among k \Rightarrow Hitting time of $\Delta_k = \{(x_i)_{i=1}^k : \exists 1 \leq i < j \leq k, x_i = x_j\}$

(larger k)

$$\mathbb{P}_\pi \left(\frac{T_k}{\mathbb{E}_\pi(T_n)} > t \right) \approx e^{-t}.$$

WHAT IS MISSING?

$$\frac{C}{n} = \sum_{k \geq 2} \frac{T_k}{n} \stackrel{?}{\sim} \sum_{k \geq 2} z_k$$

EXPONENTIALS
↑

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EXPECTATION OF T_2

HITTING OF $\Delta_k = \{(x_i)_{i \leq k} : \exists i < j, x_i = x_j\}$

EXPONENTIALS

WHAT IS MISSING?

$$\frac{C}{n} = \sum_{k \geq 2} \frac{T_k}{n} \stackrel{?}{\sim} \sum_{k \geq 2} Z_k$$

Diagram illustrating the relationship between the expectation of T_k and the hitting of Δ_k .

EXPECTATION OF T_k (indicated by a red circle around n in the first sum)

HITTING OF $\Delta_k = \{(x_i)_{i \leq k} : \exists i < j, x_i = x_j\}$ (indicated by a green circle around T_k in the first sum and a blue circle around Z_k in the second sum)

EXPONENTIALS (indicated by a blue arrow pointing to the second sum)

Problem # 1

Better estimates for tail of T_k .

WHAT IS MISSING?

$$\frac{C}{n} = \sum_{k \geq 2} \frac{T_k}{n} \stackrel{?}{\sim} \sum_{k \geq 2} Z_k$$

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Better estimates for tail of T_k .

WHAT IS MISSING?

$$\frac{C}{n} = \sum_{k \geq 2} \frac{T_k}{n} \stackrel{?}{\sim} \sum_{k \geq 2} Z_k$$

Diagram illustrating the relationship between the expectation of T_k and the hitting time Z_k .

Annotations:

- EXPONENTIALS** (blue text) with an arrow pointing to Z_k .
- EXPECTATION OF T_2** (red text) with an arrow pointing to n .
- HITTING OF** (green text) with an arrow pointing to Z_k .
- Definition:** $\Delta_k = \{(x_i)_{i \leq k} : \exists i < j, x_i = x_j\}$ (green text).

Problem # 2

Need to consider non-stationary starts.

WHAT IS MISSING?

$$\frac{C}{n} = \sum_{k \geq 2} \frac{T_k}{n} \stackrel{?}{\approx} \sum_{k \geq 2} Z_k$$

EXPECTATION OF T_2

HITTING OF $\Delta_k = \{(x_i)_{i \leq k} : \exists i < j, x_i = x_j\}$

EXPONENTIALS

Problem # 3 Must show

$$\mathbb{E}(T_n) \approx \frac{n}{\binom{k}{2}}$$

WHAT IS MISSING?

$$\frac{C}{n} = \sum_{k \geq 2} \frac{T_k}{n} \stackrel{?}{\sim} \sum_{k \geq 2} Z_k$$

EXPECTATION OF T_2

HITTING OF $\Delta_k = \{(x_i)_{i \leq k} : \exists i < j, x_i = x_j\}$

EXPONENTIALS

Problem # 4

k "big" \Rightarrow not exp.
(Big Bang)

EXPONENTIAL HITTING TIMES (NEW)

Sharper theorem.

$P =$ chain on Ω , $A \subset \Omega$ with
 π stationary; $\varepsilon = O\left(\left(\frac{t_{\text{mix}}}{\mathbb{E}_{\pi}(\tau_A)}\right)^{1/2}\right) \ll 1$

$$1) \quad \mathbb{P}_x\left(\frac{\tau_A}{\mathbb{E}_{\pi}(\tau_A)} > t\right) \leq (1 + \varepsilon) e^{-\frac{t}{1 + \varepsilon}}$$

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$P =$ chain on Ω , $A \subset \Omega$ with
 π stationary; $\varepsilon = \Omega \left(\frac{t_{\text{mix}}}{\mathbb{E}_{\pi}(\tau_A)} \right)^{1/3} \ll 1$

$$2) \quad \mathbb{P}_x \left(\frac{\tau_A}{\mathbb{E}_{\pi}(\tau_A)} > t \right) \geq (1 - \varepsilon) e^{-\frac{t}{1-\varepsilon}}$$

EXPONENTIAL HITTING TIMES (NEW)

$$\varepsilon' = \varepsilon + \mathbb{P}_x(\tau_A < \varepsilon \mathbb{E}_\pi(\tau_A))$$

Sharper theorem.

$P =$ chain on Ω , $A \subset \Omega$ with
 π stationary; $\varepsilon = \Omega\left(\left(\frac{t_{\text{mix}}}{\mathbb{E}_\pi(\tau_A)}\right)^{2/3}\right) \ll 1$

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$P = \text{chain on } \Omega$, $A \subset \Omega$ with
 π stationary; $\varepsilon = \Omega \left(\frac{t_{\text{mix}}}{\mathbb{E}_{\pi}(\tau_A)} \right)^{2/3} \ll 1$

3) ε -quantile of $\tau_A \approx \varepsilon \mathbb{E}_{\pi}(\tau_A)$

APPLICATION TO $E(T_k)$

Recall $T_k =$ Hitting time of
 $\Delta_k = \left\{ (x_i)_{i=1}^k : \exists i < j \atop x_i = x_j \right\}$

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Recall $T_k =$ Hitting time of
 $\Delta_k = \left\{ (x_i)_{i=1}^k : \exists i < j \atop x_i = x_j \right\}$

$$\Rightarrow \mathbb{P}_{\pi}(T_k \leq \varepsilon \mathbb{E}(T_2)) \lesssim \binom{k}{2} \varepsilon$$

APPLICATION TO $E(T_k)$

Recall $T_k =$ Hitting time of

$$\Delta_k = \left\{ (x_i)_{i=1}^k : \exists i < j \text{ such that } x_i = x_j \right\}$$

Also need
reverse inequality!

$$\Rightarrow \mathbb{P}_{\pi}(T_k \leq \varepsilon \mathbb{E}(T_2)) \leq \binom{k}{2} \varepsilon$$

APPLICATION TO $E(T_k)$

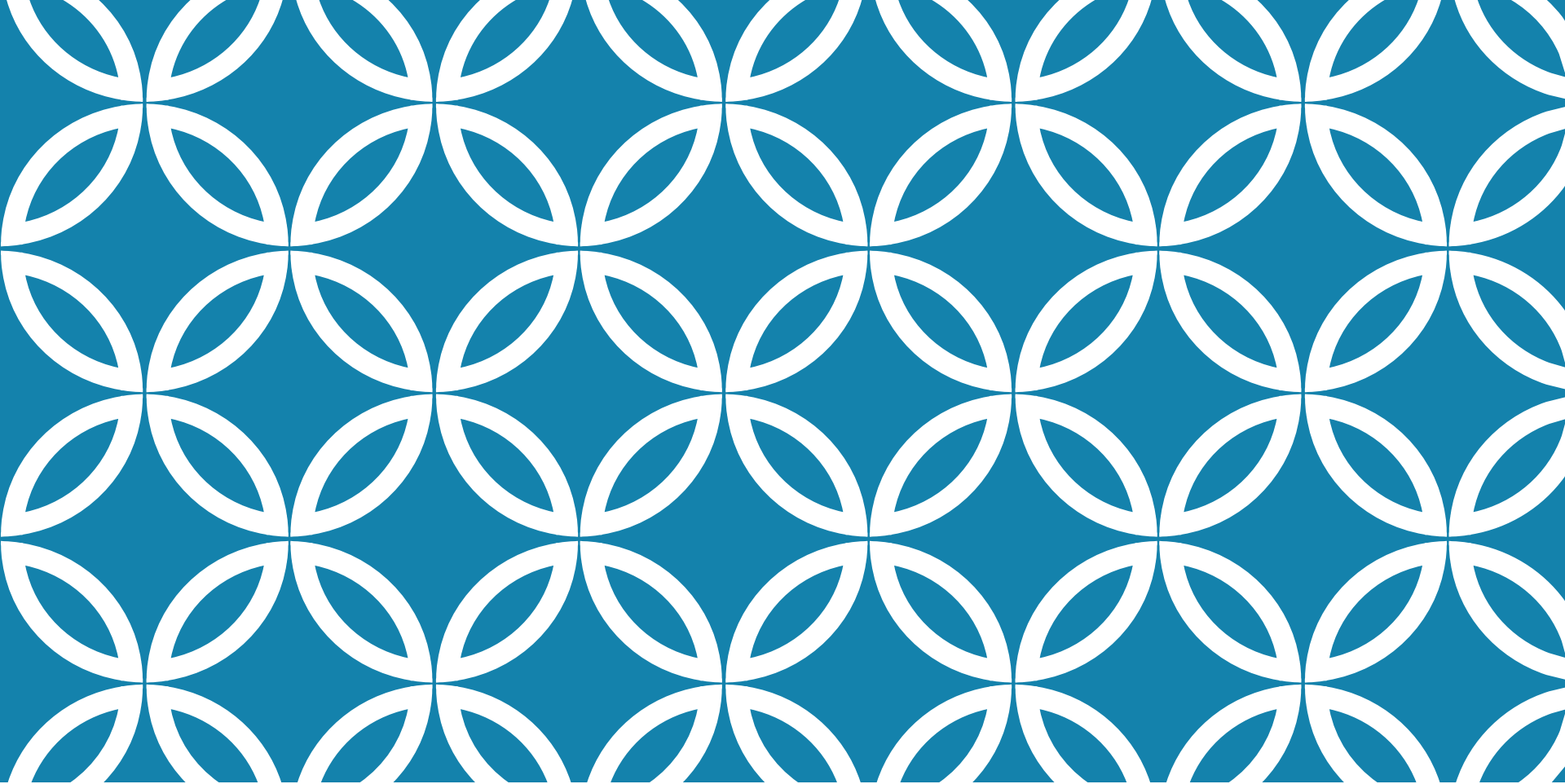
$$\mathbb{P}_\pi(T_k \leq \varepsilon \mathbb{E}(T_2)) \geq \left(\frac{k}{2}\right) \varepsilon$$

$$- O(k^4) \mathbb{P}_\pi \left(\begin{array}{l} T_2^{1,2} \leq \varepsilon \mathbb{E}(T_2), \\ T_2^{2,3} \leq \varepsilon \mathbb{E}(T_2) \end{array} \right)$$

BOUNDING CORRELATIONS (TRANSITIVE)

$$\mathbb{P}_{\pi^{\otimes 3}}(T^{1,2} \leq t, T^{2,3} \leq t) \\ \leq 2 \mathbb{P}_{\pi^{\otimes 2}}(T^{1,2} \leq t)^2$$

(on blackboard!)



THE END

Thanks for your attention.
Here is a [link to the paper](#).